

## 52.1, 2.2 Markov chains introduction

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Law of Total Probability:

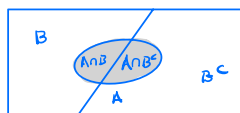
Given events  $A$  and  $B$ , with  $P(B) \neq 0$ ,

$$P(A) = P(A|B)P(B) + P(A|B^c)P(B^c).$$

More generally, if  $B_1, B_2, B_3, \dots, B_n$  are disjoint events that make up the whole sample space, with  $P(B_1), \dots, P(B_n) \neq 0$ ,

$$P(A) = P(A|B_1)P(B_1) + \dots + P(A|B_n)P(B_n).$$

Idea of proof



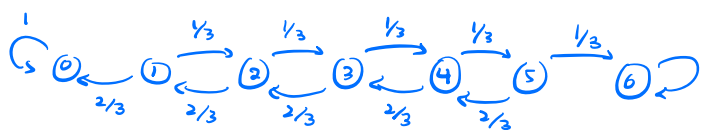
$$P(A) = P(A \cap B) + P(A \cap B^c)$$

$$= P(A|B)P(B) + P(A|B^c)P(B^c) \quad \text{since } P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Try the general version for yourself.

Example A gambler plays a game where they win \$1 each round with probability  $\frac{1}{3}$  and lose \$1 with probability  $\frac{2}{3}$ . They stop playing if their net winnings ever reach \$0 or \$6.

They start the game with \$1, 2, 3, 4, or 5 uniformly at random. Draw the transition state diagram and find the probability their net winnings are \$3 after 1 round.



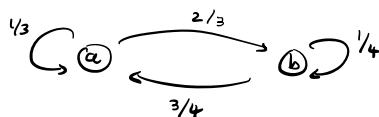
$$\begin{aligned}
 P(\bar{X}_1 = 3) &= \sum_{k=1}^5 P(\bar{X}_1 = 3 | \bar{X}_0 = k) P(\bar{X}_0 = k) \\
 &= P(\bar{X}_1 = 3 | \bar{X}_0 = 4) P(\bar{X}_0 = 4) + P(\bar{X}_1 = 3 | \bar{X}_0 = 2) P(\bar{X}_0 = 2) \\
 &= (2/3)(1/5) + (1/3)(1/5) = 1/5
 \end{aligned}$$

Conditional Law of Total Probability

$$P(A|C) = P(A|B^cC)P(B^c|C) + P(A|BC)P(B|C)$$

More generally,  $P(A|C) = P(A|B_1C)P(B_1|C) + \dots + P(A|B_nC)P(B_n|C)$

Example Consider the Markov chain with transition state diagram below.



$$\begin{aligned}
 \text{Find } P(\bar{X}_2 = a | \bar{X}_0 = a) & \quad \text{for a Markov chain,} \\
 & \quad \text{older part of history is irrelevant} \\
 &= P(\bar{X}_2 = a | \bar{X}_1 = a, \bar{X}_0 = a) P(\bar{X}_1 = a | \bar{X}_0 = a) \\
 & \quad + P(\bar{X}_2 = a | \bar{X}_1 = b, \bar{X}_0 = a) P(\bar{X}_1 = b | \bar{X}_0 = a) \\
 &= (1/3)^2 + (3/4)(2/3)
 \end{aligned}$$

Def Let  $S$  be a discrete set (so  $S$  is finite or countably infinite). A discrete-time, discrete-state Markov chain is a sequence of random variables  $(\bar{X}_0, \bar{X}_1, \bar{X}_2, \dots)$  with the property

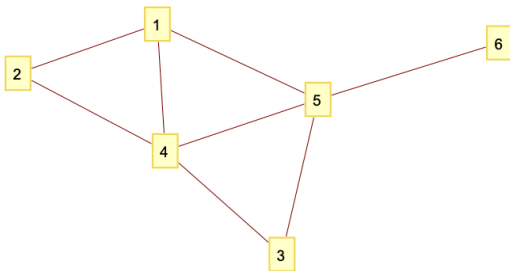
$$P(\bar{X}_{n+1} = j \mid \bar{X}_0 = x_0, \bar{X}_1 = x_1, \dots, \bar{X}_{n-1} = x_{n-1}, \bar{X}_n = i) \\ = P(\bar{X}_{n+1} = j \mid \bar{X}_n = i)$$

for any time  $n \geq 0$  and any states  $i, j, x_0, \dots, x_{n-1} \in S$ .

This is called the Markov property. A Markov chain is called time-homogeneous if

$$P(\bar{X}_{n+1} = j \mid \bar{X}_n = i) = P(\bar{X}_1 = j \mid \bar{X}_0 = i)$$

for all times  $n \geq 0$ .



**Problem 1.** Try the following exercises using the graph shown above.

- Find the transition matrix  $P$ .
- What do the following conditional probabilities mean in words? How many time units elapse? Between which states do you transition? What does the notion of *time-homogeneity* tell us about parts b and c?
  - $P(X_2 = 3 \mid X_0 = 1)$
  - $P(X_7 = 5 \mid X_4 = 4)$
  - $P(X_{50} = 5 \mid X_{40} = 2)$
- Try computing the probability in part 1 above using the Conditional Law of Total Probability. Wait until after doing Problem 2 before you think about the other probabilities.

$$P = \begin{matrix} & \begin{matrix} 1 & 2 & 3 & 4 & 5 & 6 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 3 \\ 4 \\ 5 \\ 6 \end{matrix} & \begin{bmatrix} 0 & 1/3 & 0 & 1/3 & 1/3 & 0 \\ 1/2 & 0 & 0 & 1/2 & 0 & 0 \\ 0 & 0 & 0 & 1/2 & 1/2 & 0 \\ 1/4 & 1/4 & 1/4 & 0 & 1/4 & 0 \\ 1/4 & 0 & 1/4 & 1/4 & 0 & 1/4 \\ 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix} \end{matrix}$$

(b) (1) prob. of transitioning from 1 to 3 in 2 time units

(2) time-homogeneity  $\Rightarrow P(\bar{X}_7=5 | \bar{X}_4=4) = P(\bar{X}_3=5 | \bar{X}_0=4)$

prob. of transitioning from 4 to 5 in 3 time units

(3) time-homogeneity  $\Rightarrow P(\bar{X}_{50}=5 | \bar{X}_{40}=2) = P(\bar{X}_{10}=5 | \bar{X}_0=2)$

prob. of transitioning from 2 to 5 in 10 time units

$$\begin{aligned}
 \textcircled{c} \quad P(\bar{X}_2=3 | \bar{X}_0=1) &= \sum_{k=1}^6 P(\bar{X}_2=3 | \bar{X}_1=k, \bar{X}_0=1) P(\bar{X}_1=k | \bar{X}_0=1) \\
 &= \sum_{k=1}^6 P(\bar{X}_2=3 | \bar{X}_1=k) P(\bar{X}_1=k | \bar{X}_0=1) \\
 &= \sum_{k=1}^6 P_{k3} P_{1k} \\
 &= \sum_{k=1}^6 P_{1k} P_{k3} \\
 &= \text{dot product of row 1 and} \\
 &\quad \text{column 3 of the transition matrix } P \\
 &= (P^2)_{13} \quad (\text{ie. the } (1,3) \text{ entry of } P^2) \\
 &= (0)(0) + (1/3)(0) + (0)(0) \\
 &\quad + (1/3)(1/4) + (1/3)(1/4) + (0)(0) \\
 &= 1/6
 \end{aligned}$$

**Problem 2.** Let's think more generally now. Suppose we're working with a Markov chain whose state space is  $\mathcal{S} = \{1, 2, \dots, m\}$ . Its transition matrix  $P$  is an  $m \times m$  matrix and its  $ij$ -entry is  $P_{ij} = P(X_1 = j | X_0 = i)$ . This is all given to us. Try using the Conditional Law of Total Probability to write an expression for

$$P(X_2 = j | X_0 = i),$$

called the 2-step transition probability, in terms of the entries of the transition matrix  $P$ . Your expression should be a summation and its terms should be written using entries of the matrix  $P$ . Do you recognize this expression as something related to matrix algebra? If you figure these questions out, try thinking about  $P(X_3 = j | X_0 = i)$ , called the 3-step transition probability.

$$P(\bar{X}_2 = j | \bar{X}_0 = i) = (P^2)_{ij}$$

(prove this)

$$P(\bar{X}_3 = j | \bar{X}_0 = i) = (P^3)_{ij}$$